

Genetic Associations Between Common Medication Use and Osteoporosis-Related Outcomes: A Mendelian Randomization Study

Keywords

Osteoporosis, Drug-Related Side Effects and Adverse Reactions, Mendelian Randomization Analysis, Thyroid Hormones, Anilides, Antidiabetic Agents

Abstract

Introduction

Osteoporosis pathogenesis involves a complex interplay of genetic predisposition and medication exposure. Although long-term medication use contributes to bone fragility, which is distinct from osteomalacia, genetic evidence linking common medication-use traits to osteoporosis remains limited.

Material and methods

Two-sample Mendelian randomization (MR) analyses utilized GWAS data for 22 medication-use traits (UK Biobank) and osteoporosis outcomes (FinnGen R12). Multivariable MR (MVMR) was employed to adjust for confounding by underlying disease indications. Primary estimates were derived using the inverse variance-weighted (IVW) method, supported by comprehensive sensitivity analyses including MR-Egger, weighted median, and MR-PRESSO to evaluate heterogeneity and pleiotropy.

Results

Univariable MR identified seven significant associations: anilide use with increased osteoporosis risk (OR=1.25, FDR=0.0201); immunosuppressant use with higher risks of osteoporosis with pathological fracture (OR=1.16, FDR=0.0074) and drug-induced osteoporosis (OR=1.36, FDR=0.0168); thyroid preparations with increased osteoporosis (OR=1.08, FDR=0.0016), drug-induced osteoporosis, and osteoporosis with pathological fracture; and an inverse association between antidiabetic drug use and osteoporosis (OR=0.88, FDR=0.0002). In MVMR, associations for anilide and thyroid preparation use remained significant after adjusting for back pain and hypothyroidism, respectively. However, the antidiabetic signal reversed after type 2 diabetes adjustment, indicating the observed benefit was likely confounded by the underlying metabolic condition rather than a direct pharmacological effect.

Conclusions

Genetic evidence supports that anilide and thyroid hormone use may independently increase osteoporosis risk, while the protection from antidiabetic therapy is confounded by diabetes status. Medication-related osteoporosis should be considered in long-term pharmacologic management.

1 **Genetic Associations Between Common Medication Use and Osteoporosis-Related Outcomes:**
2 **A Mendelian Randomization Study**

3

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15 **Abstract**

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19 osteoporosis remains limited.

20 **Methods:** Two-sample Mendelian randomization (MR) analyses utilized GWAS data for 22
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31 inverse association between antidiabetic drug use and osteoporosis (OR=0.88, FDR=0.0002). In
32 MVMR, associations for anilide and thyroid preparation use remained significant after adjusting
33 for back pain and hypothyroidism, respectively. However, the antidiabetic signal reversed after
34 type 2 diabetes adjustment, indicating the observed benefit was likely confounded by the
35 underlying metabolic condition rather than a direct pharmacological effect.

36 **Conclusion:** Genetic evidence supports that anilide and thyroid hormone use may independently
37 increase osteoporosis risk, while the protection from antidiabetic therapy is confounded by

38 diabetes status. Medication-related osteoporosis should be considered in long-term pharmacologic
39 management.

40

41 **Key words:** Mendelian Randomization Analysis, Osteoporosis, Drug-Related Side Effects and
42 Adverse Reactions, Anilides, Thyroid Hormones, Antidiabetic Agents, Genetic Association,
43 Pharmacology, Management

44

45 **Highlights:**

- 46 • Anilide and thyroid hormone use are genetically associated with increased osteoporosis risk.
- 47 • Antidiabetic drug use shows a protective association only before adjusting for type 2 diabetes.
- 48 • Immunosuppressant-related osteoporosis risk is largely driven by underlying rheumatoid arthritis.
- 49 • Medication-related bone fragility warrants regular monitoring in long-term users.

50 Introduction

51 Osteoporosis is a chronic metabolic bone disorder characterized by reduced bone mass and
52 deterioration of bone microarchitecture, leading to increased fragility and fracture risk[1]. Its
53 clinical manifestations primarily include pathological and fragility fractures at typical sites such
54 as the femur, lumbar spine, and pelvis[2]. Epidemiologically, vertebral fractures are among the
55 most common, with an imaging-based prevalence of approximately 30% in women and 19% in
56 men over age 50, typically assessed via morphometry on lateral spine imaging [2]. Globally,
57 osteoporosis affected approximately 41.5 million individuals in 2019 and is projected to rise to
58 263 million by 2034, representing a major and growing public health burden, particularly among
59 individuals aged ≥ 50 years[3]. Although age, hormonal imbalance, and metabolic dysregulation
60 are established contributors, emerging evidence highlights the role of long-term medication
61 exposure in accelerating bone loss and fracture susceptibility[4]. Given the widespread and often
62 chronic use of medications in older adults, elucidating their potential associations with
63 osteoporosis risk is critical for improving risk stratification and precision prevention [5].

64 A number of commonly used medications have been implicated in adversely affecting bone health
65 [6]. For instance, long-term glucocorticoid therapy is the most frequent cause of iatrogenic
66 osteoporosis, with fractures occurring in up to 30–50% of patients under chronic use[7, 8]. It
67 induces rapid bone loss, especially within the first months of treatment, through pathways
68 including suppression of osteoblast activity, promotion of osteoclastogenesis, and reduced
69 intestinal calcium absorption [7]. Thyroid hormone therapy (e.g. levothyroxine) has been
70 associated with reduced bone mass and altered bone metabolism, primarily by stimulating
71 osteoclast activity and accelerating the bone remodeling process [9]. Although these associations
72 have been frequently observed in epidemiological studies, they are often influenced by

73 confounding from underlying diseases and clinical treatment indications [10]. Mendelian
74 randomization (MR) uses genetic variants as instrumental variables (IVs) to estimate potential
75 causal effects while minimizing bias from environmental and clinical factors, providing a reliable
76 approach to disentangle complex relationships[11]. Previous MR studies have reported
77 heterogeneous effects of antidiabetic, lipid-lowering, thyroid-related, and reproductive factors on
78 osteoporosis risk or bone mineral density (BMD)[12-15]. However, these studies have focused on
79 single drug classes and outcomes [16]. This limitation necessitates a comprehensive genetic
80 assessment spanning multiple medication-use and skeletal phenotypes, encompassing osteoporosis,
81 pathological fractures, and other related skeletal outcomes, to clarify medication-related
82 associations and guide safer strategies for osteoporosis prevention and management [10], such as
83 providing a genetic basis for personalized medication decisions and reinforcing bone health
84 monitoring in long-term users of implicated drugs.

85 Therefore, to overcome the limitations of previous piecemeal analyses, we conducted a
86 comprehensive, systematic genetic investigation. This study aimed to simultaneously evaluate the
87 genetic associations of a broad spectrum of common medication classes with multiple
88 osteoporosis-related skeletal outcomes, and to employ multivariable MR (MVMR) to disentangle
89 direct pharmacologic effects from confounding by underlying disease indications. By integrating
90 this dual analytical framework across diverse drug categories and phenotypes, our work provides
91 novel, genetically-informed evidence on medication-related osteoporosis risk, with the ultimate
92 goal of guiding safer prescribing practices and targeted bone health monitoring.

93

94 **Materials and methods**

95 **Study design**

96 This study followed a pre-defined, stepwise analytical plan to assess the causal role of medication
97 use in osteoporosis. The methodology adhered to the STROBE-MR reporting guidelines [17], and
98 the overall analytical workflow is conceptually outlined in **Figure 1**. The detailed sequence of
99 steps is as follows:

100 **Step 1. Data Acquisition and Instrumental Variable (IV) Selection.** We obtained summary-level
101 genetic data from large-scale genome-wide association studies (GWAS). This included: (i) 22
102 medication-use traits as exposures (UK Biobank), (ii) osteoporosis and related fracture outcomes
103 (FinnGen R12), and (iii) genetic predispositions to key confounding diseases (e.g., type 2 diabetes,
104 back pain; Million Veteran Program). For each trait, independent single-nucleotide polymorphisms
105 (SNPs) meeting genome-wide significance ($P < 5 \times 10^{-8}$) and linkage disequilibrium ($R^2 < 0.001$)
106 criteria were selected as IVs, ensuring adherence to the three core MR assumptions [18].

107 **Step 2. Univariable Mendelian Randomization Analysis.** Using the selected IVs, we first
108 performed two-sample MR analyses for each exposure-outcome pair to estimate the total genetic
109 association. The primary causal estimate was derived using the inverse-variance weighted (IVW)
110 method, with sensitivity estimates from MR-Egger, weighted median, and weighted mode methods.

111 **Step 3. Multivariable MR for Confounding Adjustment.** To disentangle the direct effect of
112 medication use from the effect of the underlying disease indication, we performed multivariable
113 MR (MVMR). For each significant or candidate medication, such as anilides, thyroid drugs, we
114 included the genetic instruments for both the medication and its corresponding indication disease
115 such as back pain, hypothyroidism in a joint model using the multivariable IVW method.

116 **Step 4. Sensitivity and Robustness Validation.** The validity of all significant associations from
117 Steps 2 and 3 was rigorously tested. This involved: (i) assessing heterogeneity among IVs using
118 Cochran's Q statistic; (ii) testing for horizontal pleiotropy using MR-Egger intercept and MR-

119 PRESSO global tests; and (iii) conducting leave-one-out analyses to ensure results were not driven
120 by a single influential SNP.

121 Data sources

122 GWAS data for drug exposure traits were obtained from a large-scale UK Biobank analysis by Wu
123 et al., who performed GWAS of 23 self-reported medication-use categories in approximately
124 320,000 participants of European ancestry to identify genetic variants associated with medication-
125 taking behaviors[19]. After excluding drugs affecting bone structure and mineralization use
126 measurement (GCST007935) because of its direct biological relevance to osteoporosis outcomes,
127 22 medication-use traits were retained for analysis.

128 Outcome data for osteoporosis and related phenotypes were derived from the FinnGen R12 cohort,
129 including osteoporosis (M13_OSTEOPOROSIS; 10,461 cases and 473,264 controls), osteoporosis
130 with pathological fracture (OSTEOPOROSIS_FRACTURE_FG; 2,429 cases and 374,317
131 controls), fracture of the femur (ST19_FRACT_FEMUR; 12,229 cases and 476,877 controls),
132 fracture of the lumbar spine and pelvis (ST19_FRACT_LUMBAR_SPINE_PELVIS; 9,061 cases
133 and 481,639 controls), coxarthrosis (M13_ARTHTROSIS_COX; 30,802 cases and 315,115
134 controls), and drug-induced osteoporosis (DRUGADVERS_OSTEOPO; 389 cases and 499,959
135 controls). Diagnostic definitions for the outcomes were based on national registry data using
136 standardized ICD-8, ICD-9, and ICD-10 codes.

137 To account for potential confounding from underlying diseases, MVMR was performed using
138 GWAS summary data from the Million Veteran Program (MVP). The included traits were type 2
139 diabetes (T2D; Phe_250_2; 154,194 cases and 278,454 controls), hypothyroidism (Phe_244;
140 62,814 cases and 378,321 controls), rheumatoid arthritis (RA; Phe_714_1; 10,479 cases and

141 436,161 controls), and back pain (Phe_760; 218,071 cases and 185,068 controls), corresponding
142 respectively to antidiabetic, thyroid, immunosuppressant, and anilide exposures.

143 All included traits are summarized in **Table 1**. This study used only publicly available, de-
144 identified GWAS summary statistics. All original studies received ethical approval from their
145 respective institutional review boards, and informed consent was obtained from participants. No
146 additional ethical approval or consent was required for the present analyses. **The original data used**
147 **in this study, including the summary statistics and the full code used for the MR and MVMR**
148 **analyses, have been uploaded to the Zenodo database with the DOI:**
149 **[<https://doi.org/10.5281/zenodo.17970613>]. The data can be accessed via the following link:**
150 **[<https://doi.org/10.5281/zenodo.17970613>].**

151 **Selection of IVs**

152 SNPs associated with each exposure were selected at a genome-wide significance threshold of P
153 $< 5 \times 10^{-8}$. For exposures yielding fewer than four instruments, a relaxed threshold of $P < 5 \times 10^{-6}$
154 was applied[20], including peptic ulcer and GORD drugs (GCST007922), vasodilators used in
155 cardiac diseases (GCST007925), antihypertensives (GCST007926), immunosuppressants
156 (GCST007933), opioids (GCST007936), anilides (GCST007938), and antidepressants
157 (GCST007940). SNPs with a minor allele frequency (MAF) ≤ 0.01 were excluded to avoid bias
158 from rare variants. Linkage disequilibrium (LD) pruning was conducted using an $R^2 < 0.001$ and
159 a window size of 10,000 kb to retain independent variants. When a selected SNP was unavailable
160 in the outcome dataset, a proxy SNP in high LD ($R^2 > 0.8$) was substituted. The strength of each
161 IV was evaluated using the F statistic, calculated as $F = R^2 \times (N - 2) / (1 - R^2)$. R^2 represents the
162 proportion of exposure variance explained by the SNP. Instruments with F values ≤ 10 were
163 excluded to minimize weak instrument bias[21].

164 Outlier removal

165 The MR-PRESSO framework was applied for outlier detection and correction, including a global
166 test for overall pleiotropy, an outlier test to identify and remove heterogeneous variants, and a
167 distortion test to assess whether causal estimates changed significantly after correction[22].
168 Additionally, an LOO analysis was conducted for each exposure–outcome pair to evaluate the
169 influence of individual SNPs on the overall MR estimate. For exposure–outcome pairs that still
170 exhibited heterogeneity or horizontal pleiotropy, the LOO procedure was further performed
171 iteratively by excluding the most influential variant in each round, repeating the MR-PRESSO,
172 MR-Egger heterogeneity (Q-test), and MR-Egger intercept tests until all three criteria were
173 simultaneously satisfied (MR-PRESSO global test $P > 0.05$, MR-Egger Q-test $P > 0.05$, and MR-
174 Egger intercept $P > 0.05$).

175 Univariable MR analysis

176 Putative causal effects between each medication-use trait and osteoporosis-related outcomes were
177 estimated using four complementary MR models: inverse variance–weighted (IVW), MR-Egger,
178 weighted median, and weighted mode. The IVW method served as the primary analysis to provide
179 pooled causal estimates under the assumption of no horizontal pleiotropy. The remaining three
180 approaches were applied to assess the stability and consistency of the results[23]. Multiple testing
181 correction was applied to IVW-derived P values using the Benjamini–Hochberg method. A false
182 discovery rate (FDR) < 0.05 was considered statistically significant.

183 Sensitivity analysis

184 To assess the validity of IVs and the robustness of causal estimates, a series of sensitivity analyses
185 was performed. Heterogeneity among SNPs was evaluated using Cochran’s Q statistic under the
186 IVW model ($P < 0.05$ indicating significant heterogeneity). Horizontal pleiotropy was examined

187 using the MR-Egger intercept test ($P < 0.05$ suggesting directional pleiotropy) and further assessed
188 through the Mendelian Randomization Pleiotropy RESidual Sum and Outlier (MR-PRESSO)
189 global test [24, 25].

190 **MVMR**

191 To control for confounding by underlying disease indications, multivariable MR was performed
192 using genome-wide significant SNPs ($P < 5 \times 10^{-8}$, $MAF > 0.01$) associated with both drug-use
193 exposures and corresponding disease traits [26]. LD pruning was independently conducted for each
194 trait using $R^2 < 0.001$ and a 10,000 kb window to remove correlated variants. The pruned SNP sets
195 from all exposures and confounders were merged, deduplicated, and re-pruned under the same
196 parameters to generate a pool of mutually independent candidate variants. From this joint set, SNPs
197 were extracted separately for the exposure and confounder datasets. To ensure strand consistency,
198 exposure datasets were treated as exposures and confounder datasets as outcomes during
199 harmonization. Effect alleles were corrected for non-palindromic variants, and all palindromic
200 SNPs were excluded. After harmonization, only SNPs present in both exposure and confounder
201 datasets were retained as final IVs. MVMR analyses were conducted using the multivariable IVW
202 approach to estimate the effects of each exposure while accounting for correlated traits [27].

203 **Statistical analysis and image generation**

204 All statistical analyses were conducted using R software (version 4.3.0) with the “TwoSampleMR”
205 package. Scatter, Forest, and funnel plots were generated to visualize MR results.

206

207 **Results**

208 **IV selection**

209 Osteoporosis is often clinically silent until a fracture occurs. Its most common manifestations are
210 fragility fractures resulting from minimal trauma, typically at the spine, hip, or wrist. Vertebral
211 fractures can lead to back pain, height loss, and kyphosis, while hip fractures are associated with
212 significant morbidity and mortality; large-scale studies have recorded tens of thousands of deaths
213 attributable to such fractures [28]. The outcomes analyzed in this study directly capture these
214 critical clinical endpoints. In postmenopausal women, the accelerated bone loss that primarily
215 underlies osteoporosis is driven by estrogen deficiency. This hormonal change leads to increased
216 bone resorption, resulting in rapid reduction of bone mineral density and deterioration of trabecular
217 microstructure [29], thereby increasing skeletal fragility. Our analysis of osteoporosis-related
218 genetic associations directly addresses this significant clinical and public health burden.

219 After harmonization and quality control, independent genome-wide significant variants were
220 selected for 22 medication-related exposures. Outlier detection and sensitivity analyses were
221 performed for all exposure–outcome pairs. To ensure robust causal estimates, SNPs with
222 significant horizontal pleiotropy or undue influence were excluded through a two-step procedure:
223 (1) initial outlier detection using the MR-PRESSO framework, followed by (2) an iterative leave-
224 one-out (LOO) analysis for pairs where heterogeneity persisted. The LOO iteration continued until
225 the estimates satisfied pre-specified criteria for heterogeneity and pleiotropy (all $P > 0.05$). The list
226 of excluded SNPs is detailed in **Table S1**. After exclusion, all retained IVs satisfied the predefined
227 convergence criteria. The final number of SNPs per exposure–outcome pair ranged from 4
228 (antihistamine–coxarthrosis) to 142 (renin–angiotensin system agents–drug-induced osteoporosis).

229 **MR analyses of medication use and osteoporosis-related skeletal outcomes**

230 Using the IVW method, significant associations ($FDR < 0.05$) were observed for anilide use with
231 elevated osteoporosis risk ($OR = 1.25$, 95% CI: 1.09–1.44, $FDR = 0.0201$); immunosuppressant

232 use with higher risks of osteoporosis with pathological fracture (OR = 1.16, 95% CI: 1.07–1.26,
233 FDR = 0.0074) and drug-induced osteoporosis (OR = 1.36, 95% CI: 1.13–1.64, FDR = 0.0168);
234 thyroid preparations with increased risks of osteoporosis (OR = 1.08, 95% CI: 1.04–1.12, FDR =
235 0.0016), drug-induced osteoporosis (OR = 1.37, 95% CI: 1.17–1.61, FDR = 0.0021), and
236 osteoporosis with pathological fracture (OR = 1.12, 95% CI: 1.05–1.19, FDR = 0.0158).
237 Conversely, antidiabetic drug use was inversely associated with osteoporosis (OR = 0.88, 95% CI:
238 0.84–0.93, FDR = 0.0002). Several additional IVW associations reached nominal significance (P
239 < 0.05 but $FDR \geq 0.05$), including anilides, glucocorticoids, HMG-CoA reductase inhibitors,
240 immunosuppressants, and inhalant adrenergics with osteoporosis and related skeletal disorders
241 (**Table 2**). Results from supplementary MR methods (MR-Egger, weighted median, and weighted
242 mode) were largely consistent in direction and magnitude with the IVW estimates (**Table S2**). The
243 scatter (**Figure 2**) and forest plots (**Figure S1**) for FDR-significant associations demonstrated
244 consistent SNP-specific effects and overall MR estimates across methods. The corresponding
245 funnel and LOO plots showed symmetric distributions and the absence of influential single
246 variants (**Figure S2–S3**).

247 **Sensitivity analyses**

248 Sensitivity analyses were conducted to assess heterogeneity and horizontal pleiotropy for all
249 exposure–outcome pairs after outlier correction. As shown in **Table S3**, Cochran’s Q test indicated
250 no significant heterogeneity ($P > 0.05$), and the MR-Egger intercept test showed no evidence of
251 directional pleiotropy ($P > 0.05$). Additionally, the MR-PRESSO global test detected no residual
252 horizontal pleiotropy ($P > 0.05$) in any exposure–outcome combination (**Table S4**). These findings
253 indicate that the IVs used in the MR analyses were generally valid and that the estimated
254 associations were unlikely to be biased by heterogeneity or unbalanced pleiotropy.

255 **MVMR analysis**

256 MVMR was conducted to determine whether the associations identified in univariable MR
257 persisted after adjusting for underlying disease indications. As shown in **Table 3**, the associations
258 of anilide use with elevated osteoporosis risk (OR = 1.25, 95% CI: 1.06–1.48, P = 0.0084) and
259 thyroid preparation use with increased risks of osteoporosis (OR = 1.14, 95% CI: 1.08–1.21, P =
260 1.3×10^{-5}) and osteoporosis with pathological fracture (OR = 1.17, 95% CI: 1.05–1.31, P = 0.004)
261 remained significant after controlling for back pain and hypothyroidism, respectively. In contrast,
262 the previously protective association between antidiabetic drug use and osteoporosis in univariable
263 MR reversed after adjustment for T2D (OR = 1.13, 95% CI: 1.03–1.25, P = 0.008), indicating that
264 the benefit was likely driven by disease-related confounding. Similarly, the associations of thyroid
265 preparations and immunosuppressants with drug-induced osteoporosis, and immunosuppressant
266 use with osteoporosis with pathological fracture, lost significance after controlling for
267 hypothyroidism or RA, suggesting that these relationships were largely attributable to the
268 underlying disease indications rather than the medications themselves.

269

270 **Discussion**

271 In this study, MR and MVMR analyses identified independent associations of anilide use with
272 increased osteoporosis risk and thyroid preparations with higher risks of osteoporosis and
273 osteoporosis with pathological fracture [30, 31]. The inverse association between antidiabetic drug
274 use and osteoporosis in univariable MR was reversed after adjusting for T2D, reflecting
275 confounding by indication [32]. Similarly, associations of thyroid preparations and
276 immunosuppressants with drug-induced osteoporosis lost significance after controlling for
277 hypothyroidism or RA, suggesting disease-related rather than pharmacologic effects [33]. These

278 findings indicate that prolonged use of anilide-containing analgesics, thyroid hormone
279 replacement, and antidiabetic medications may increase bone fragility, warranting regular bone
280 density monitoring and timely preventive intervention in long-term users.

281 By integrating univariable and multivariable MR approaches [34], this study revealed medication-
282 related skeletal effects while accounting for confounding by underlying disease indications. The
283 consistent associations of anilide and thyroid hormone use with increased osteoporosis risk,
284 together with the reversal of the antidiabetic drug association and the loss of significance for
285 immunosuppressant use after adjustment, indicate that the applied framework effectively
286 differentiated pharmacologic from disease-related influences. This dual-level analysis enhances
287 the reliability of genetically inferred associations and underscores the methodological value of
288 genetically predicted medication-use traits for evaluating drug safety at the population scale.

289 Anilides, primarily paracetamol, are among the most commonly used analgesics [19, 35]. While
290 epidemiological studies have associated their regular use with increased fracture risk [36], our
291 Mendelian randomization analysis provides novel genetic evidence supporting a potential causal
292 role of anilide use in osteoporosis, independent of conventional behavioral confounders. The
293 underlying mechanisms may involve the inhibition of bone-forming osteoblast activity through
294 cyclooxygenase suppression, coupled with reduced mobility due to chronic pain[37, 38]. Therefore,
295 our findings reinforce and extend previous observational concerns by suggesting a genetic
296 predisposition to anilide use is linked to bone fragility, warranting heightened vigilance regarding
297 bone health in long-term users.

298 The observed association between thyroid preparation use and increased osteoporosis risk aligns
299 with extensive evidence that thyroid hormones accelerate bone turnover [39]. It is noteworthy that
300 vitamin D status, a key regulator of calcium homeostasis, interacts with thyroid function and may

301 modulate the skeletal impact of thyroid hormones, representing an important area for future
302 investigation. Our multivariable MR analysis extends this understanding by demonstrating that the
303 risk for primary osteoporosis and pathological fracture remains significant even after accounting
304 for genetic predisposition to hypothyroidism. This suggests a direct, independent effect of thyroid
305 hormone excess on bone fragility, which is consistent with clinical studies linking even subclinical
306 hyperthyroidism to lower BMD and higher fracture risk [40, 41].

307 The reversal of the antidiabetic drug association after adjusting for T2D in our MVMR analysis
308 highlights profound confounding by indication. Although T2D patients may have altered bone
309 material strength [42] and some drugs like metformin may appear beneficial in unadjusted analyses
310 [43, 44], our findings clarify that any observed benefit likely reflects improved metabolic control
311 rather than a direct skeletal effect. This underscores that the relationship between antidiabetic
312 therapy and bone health is complex and dominated by the underlying metabolic state.

313 RA is a chronic autoimmune disease with an increasing global burden, which induces systemic
314 inflammation and cytokine release that inhibit bone formation [45]. For immunosuppressants, the
315 attenuation of associations after controlling for RA in MVMR supports a disease-related rather
316 than pharmacologic effect. This aligns with the established view that these inflammatory processes
317 are primary drivers of bone loss, effects which may outweigh direct drug toxicity [46]. Our results
318 thus suggest that the osteoporosis risk in these patients largely reflects the underlying inflammatory
319 disease burden.

320 **Limitations**

321 Several limitations of this study should be considered. First, the generalizability of our findings
322 may be limited as all genetic data were from European-ancestry populations. Second, medication-
323 use exposures were based on self-report, lacking dosage and duration information, which prevents

324 dose-response assessment and may introduce measurement error. Third, as an MR study, our
325 estimates reflect the effect of genetic predisposition to long-term medication use, not the effect of
326 initiating a specific pharmacologic intervention. Fourth, despite sensitivity analyses, residual
327 confounding from pleiotropy cannot be entirely excluded. Finally, our outcomes were confined to
328 the osteoporosis phenotypes available in FinnGen; incorporating bone imaging parameters in the
329 future would deepen the biological interpretation.

330

331 **Conclusion**

332 In conclusion, this Mendelian randomization study provides genetic evidence that the
333 predisposition to use specific medications is associated with altered osteoporosis risk. We found
334 that anilide and thyroid hormone use may independently increase the risk of osteoporosis and
335 related fractures. In contrast, the apparent protective association of antidiabetic drugs was not
336 sustained after accounting for type 2 diabetes, highlighting the critical role of confounding by
337 indication. These findings underscore that medication-related skeletal fragility is an important
338 consideration in clinical practice. Specifically, they suggest that patients on long-term paracetamol
339 (anilide) or thyroid hormone therapy may benefit from proactive bone health monitoring. By
340 systematically evaluating multiple drug classes and employing multivariable MR to disentangle
341 direct effects from disease confounding, this study demonstrates a robust genetic epidemiological
342 framework for post-marketing drug safety assessment. Future research incorporating diverse
343 populations, precise prescription data, and detailed bone imaging phenotypes will be valuable to
344 translate these genetic insights into personalized clinical strategies for osteoporosis prevention.

345 Declarations**346 Ethics approval and consent to participate**

347 This study used only publicly available, de-identified GWAS summary statistics. All original
348 studies received ethical approval from their respective institutional review boards, and informed
349 consent was obtained from participants. No additional ethical approval or consent was required for
350 the present analyses.

351 Consent for publication

352 Not applicable.

353 Availability of data and materials

354 All data generated or analyzed during this study are included in this published article and its
355 supplementary information file.

356 Competing interests

357 The authors declare that they have no competing interests.

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362 Authors' contributions

363 Yuxuan Wu and Xi Chen carried out the studies, participated in collecting data, and drafted the
364 manuscript. Xue Chen and Ran Hao performed the statistical analysis and participated in its design.
365 Zhicong Wang participated in acquisition, analysis, or interpretation of data and draft the
366 manuscript. All authors read and approved the final manuscript.

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368 Not applicable

369 **Abbreviations**

370 3-Hydroxy-3-methylglutaryl-coenzyme A, HMG-CoA

371 Advanced glycation end-products, AGEs

372 Anatomical Therapeutic Chemical Classification System, ATC

373 Bone mineral density, BMD

374 Coefficient of determination, R^2

375 Confidence interval, CI

376 Cyclooxygenase, COX

377 False discovery rate, FDR

378 Gastro-oesophageal reflux disease, GORD

379 Genome-wide association study, GWAS

380 GWAS Catalog Study ID, GCST

381 Instrumental variable, IV

382 International Classification of Diseases, ICD

383 Inverse variance-weighted, IVW

384 Leave-one-out, LOO

385 Linkage disequilibrium, LD

386 Mendelian Randomization Pleiotropy RESidual Sum and Outlier, MR-PRESSO

387 Mendelian randomization, MR

388 Minor allele frequency, MAF

389 Million Veteran Program, MVP

390 Multivariable Mendelian randomization, MVMR

- 391 Odds ratio, OR
- 392 Phenotype code, Phe
- 393 Prostaglandin E₂, PGE₂
- 394 Receptor activator of nuclear factor κ B ligand, RANKL
- 395 Rheumatoid arthritis, RA
- 396 Single-nucleotide polymorphism, SNP
- 397 Sodium–glucose cotransporter 2, SGLT2
- 398 Standard error, SE
- 399 Strengthening the Reporting of Observational Studies in Epidemiology using Mendelian
- 400 Randomization, STROBE-MR
- 401 Type 2 diabetes, T2D

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520 **Figure legends**

521 **Figure 1. Flowchart of the Mendelian randomization study design.** The schematic illustrates
522 the three-phase analytical workflow, encompassing instrumental variable selection and
523 harmonization, univariable and multivariable MR analyses, and sensitivity validation.

524 **Figure 2. Scatter plots of SNP-specific causal estimates for significant associations between**
525 **medication use and osteoporosis-related outcomes.** Each plot presents the association between
526 a medication-use trait and an osteoporosis outcome (FDR < 0.05). The slope of each line represents
527 the causal estimate from a different MR method. (A) Anilide use–osteoporosis, (B) Antidiabetic
528 drug use–osteoporosis, (C) Immunosuppressant use–osteoporosis with pathological fracture, (D)
529 Immunosuppressant use–drug-induced osteoporosis, (E) Thyroid preparation use–osteoporosis, (F)
530 Thyroid preparation use–drug-induced osteoporosis, and (G) Thyroid preparation use–
531 osteoporosis with pathological fracture.

532

533 **Supplementary figures**

534 **Figure S1. Forest plots showing SNP-specific and pooled MR estimates for significant**
535 **medication–osteoporosis associations.** Each horizontal line represents the effect size and 95%
536 confidence interval for a single SNP (FDR < 0.05). Panels (A–G) correspond to the same
537 exposure–outcome pairs described in Figure 2.

538 **Figure S2. Funnel plots assessing heterogeneity and directional pleiotropy for medication–**
539 **osteoporosis associations.** Each dot represents an individual SNP, plotted against its precision
540 (1/SE). The symmetrical shape of the funnels suggests the absence of directional pleiotropy. Panels
541 (A–G) correspond to associations with FDR < 0.05.

542 **Figure S3. Leave-one-out (LOO) analyses for significant medication–osteoporosis**
543 **associations.** Each point represents the MR estimate recalculated after excluding one SNP at a
544 time (FDR < 0.05). The stability of effect estimates across all iterations indicates that no single
545 variant disproportionately influenced the overall causal estimates. Panels (A–G) correspond to the
546 same exposure–outcome pairs described in Figure 2.

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Table 1. Summary of GWAS datasets for medication-use exposures, osteoporosis-related outcomes, and covariates included in the Mendelian randomization analysis

ID	Trait(s)	Cas e (n)	Contr ol (n)	Total Sample (n)	Categor y	Corresponding Exposure
GCST007922	Peptic ulcer and GORD drug use measurement			132367	Exposure	
GCST007923	Drugs used in diabetes use measurement			305913	Exposure	Type 2 diabetes (Phe_250_2)
GCST007924	Antithrombotic agent use measurement			153639	Exposure	
GCST007925	Vasodilators used in cardiac diseases use measurement			242659	Exposure	
GCST007926	Antihypertensive use measurement			152380	Exposure	
GCST007927	Beta blocking agent use measurement			224024	Exposure	
GCST007928	Diuretic use measurement			229086	Exposure	
GCST007929	Calcium channel blocker use measurement			204378	Exposure	

GCST007930	Agents acting on the renin–angiotensin system use measurement	237530	Exposure	
GCST007931	HMG CoA reductase inhibitor use measurement	290385	Exposure	
GCST007932	Thyroid preparation use measurement	305582	Exposure	Hypothyroidism (Phe_244)
GCST007933	Immunosuppressant use measurement	272602	Exposure	Rheumatoid arthritis (Phe_714_1)
GCST007934	Non-steroidal anti-inflammatory and antirheumatic product use measurement	164520	Exposure	
GCST007936	Opioid use measurement	78808	Exposure	
GCST007937	Aspirin (salicylic acid and derivatives) use measurement	112010	Exposure	
GCST007938	Anilide use measurement	179810	Exposure	Back pain (Phe_760)

GCST007939	Antimigraine preparation use measurement		119844		Exposure
GCST007940	Antidepressant use measurement		304162		Exposure
GCST007941	Inhalant adrenergic use measurement		176445		Exposure
GCST007942	Glucocorticoid use measurement		205700		Exposure
GCST007943	Antihistamine use measurement		151636		Exposure
GCST007944	Antiglaucoma preparations and miotics use measurement		100868		Exposure
M13_OSTEOPOROSIS	Osteoporosis	104	47326	483725	Outcome
		61	4		
OSTEOPOROSIS_FRACTURE_FG	Osteoporosis with pathological fracture	242	37431	376746	Outcome
		9	7		
ST19_FRACT_FEMUR	Fracture of femur	122	47687	489106	Outcome
		29	7		
ST19_FRACT_LUMBAR_PELVIS	Fracture of lumbar spine and pelvis	906	48163	490700	Outcome
R_SPINE_PELVIS		1	9		

M13_ARTHROSIS_CO	Coxarthrosis	308	31511	345917	Outcome		
X		02	5				
DRUGADVERS_OSTE	Drug-induced osteoporosis	389	49995	500348	Outcome		
OPO			9				
Phe_250_2	Type 2 diabetes	154	27845	432648	Covariate	Drugs used in	
		194	4		(MVP)	diabetes	
Phe_244	Hypothyroidism	628	37832	441135	Covariate	Thyroid	
		14	1		(MVP)	preparation	
Phe_714_1	Rheumatoid arthritis	104	43616	446640	Covariate	Immunosuppressa	
		79	1		(MVP)	nt	
Phe_760	Back pain	218	18506	403139	Covariate	Anilide	
		071	8		(MVP)		

Abbreviations: GWAS, genome-wide association study; MR, Mendelian randomization; MVP, Million Veteran Program; Phe, phenotype code from MVP; GORD, gastro-oesophageal reflux disease; HMG-CoA, 3-hydroxy-3-methylglutaryl-coenzyme A.

All exposure data were obtained from the UK Biobank medication-use GWAS by Wu et al. (Nat Med, 2019). Osteoporosis-related outcome data were derived from the FinnGen R12 release, and covariate data were obtained from the Million Veteran Program (MVP).

“Corresponding exposure” indicates the medication–disease pairing used in multivariable MR analyses to adjust for underlying indications.

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Table 2. Univariable Mendelian randomization estimates for the associations between medication use and osteoporosis-related outcomes

Exposure	Outcome	nSNP	Method	OR (95% CI)	P-value	FDR
Anilide	Osteoporosis	36	Inverse variance weighted	1.2541 (1.0895 - 1.4435)	0.0016	0.020
Drugs used in diabetes	Osteoporosis	47	Inverse variance weighted	0.8846 (0.8385 - 0.9332)	6.95×10^{-6}	2.00×10^{-4}
Immunosuppressant	Osteoporosis with pathological fracture	25	Inverse variance weighted	1.1602 (1.0681 - 1.2602)	4.00×10^{-4}	0.007
Immunosuppressant	Drug-induced osteoporosis	26	Inverse variance weighted	1.3617 (1.1294 - 1.6418)	0.0012	0.016

Thyroid preparation	Osteoporosis	107	Inverse variance weighted	1.0761	7.16×10^{-5}	0.001
				(1.0378	-	6
				1.1157)		
Thyroid preparation	Drug-induced osteoporosis	109	Inverse variance weighted	1.3730	1.00×10^{-4}	0.002
				(1.1698	-	1
				1.6115)		
Thyroid preparation	Osteoporosis with pathological fracture	109	Inverse variance weighted	1.1167	0.001	0.015
				(1.0455	-	8
				1.1928)		
Anilide	Fracture of femur	36	Inverse variance weighted	1.2084	0.006	0.060
				(1.0557	-	1
				1.3831)		
Anilide	Coxarthrosis	28	Inverse variance weighted	1.1746	0.0111	0.102
				(1.0373	-	5
				1.3300)		

Antimigraine preparation	Coxarthrosis	11	Inverse variance weighted	1.0610	0.016	0.130
				(1.0111	-	1
				1.1134)		
Antithrombotic agent	Coxarthrosis	11	Inverse variance weighted	0.8734	0.0251	0.192
				(0.7759	-	3
				0.9832)		
Drugs used in diabetes	Coxarthrosis	43	Inverse variance weighted	0.9561	0.0152	0.130
				(0.9221	-	1
				0.9914)		
Glucocorticoid	Fracture of femur	18	Inverse variance weighted	1.0972	0.0303	0.208
				(1.0089	-	9
				1.1932)		
HMG CoA reductase inhibitor	Osteoporosis	65	Inverse variance weighted	0.8879	0.0047	0.053
				(0.8177	-	5
				0.9641)		

Immunosuppressant	Fracture of femur	26	Inverse variance weighted	1.0527	0.0061	0.060
				(1.0148	-	1
				1.0920)		
Inhalant adrenergic	Osteoporosis with pathological fracture	50	Inverse variance weighted	1.1445	0.0295	0.208
				(1.0135	-	9
				1.2925)		
Inhalant adrenergic	Drug-induced osteoporosis	50	Inverse variance weighted	1.3829	0.0343	0.225
				(1.0243	-	1
				1.8669)		
Non-steroidal anti-inflammatory antirheumatic product	Osteoporosis with pathological fracture	7	Inverse variance weighted	2.0261	0.0381	0.239
				(1.0394	-	2
				3.9495)		

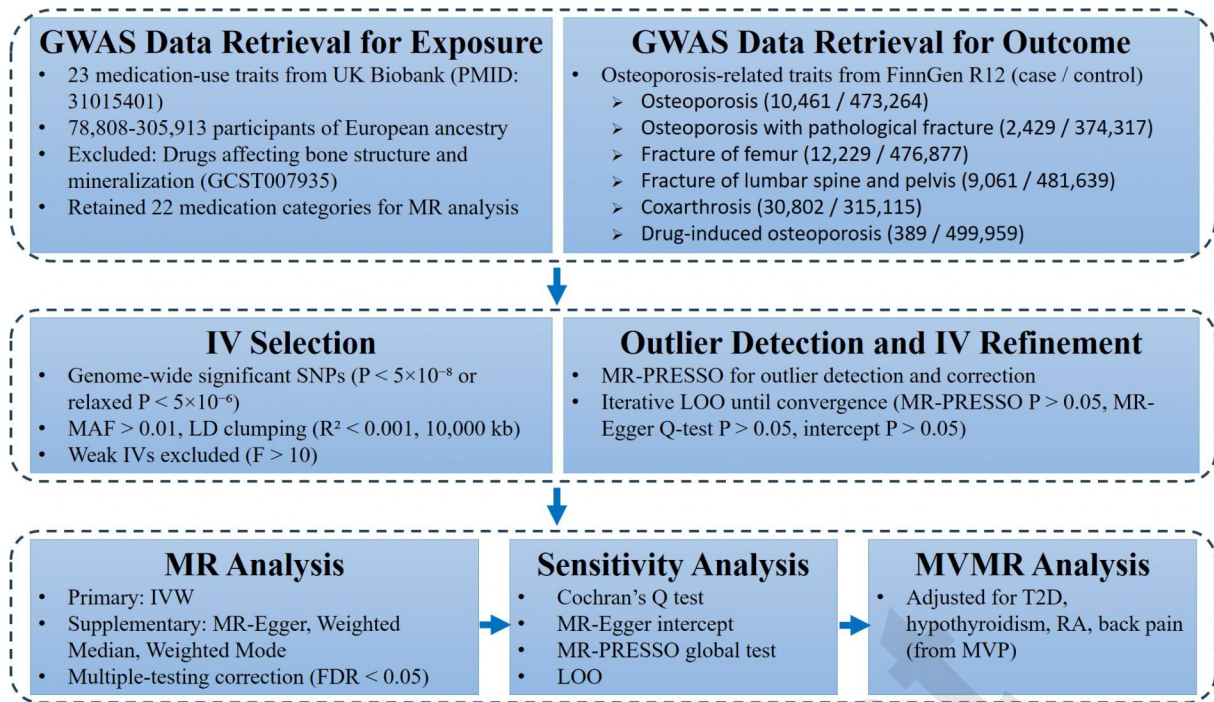
Estimates are derived from inverse variance-weighted (IVW) analyses using genome-wide significant and linkage-disequilibrium-independent SNPs as instrumental variables. Odds ratios (OR) represent the effect of genetically predicted medication use on the specified outcome. FDR, false discovery rate, was calculated using the Benjamini-Hochberg method.

Table 3. Multivariable Mendelian randomization (MVMR) estimates for the associations between medication use and osteoporosis-related outcomes after adjustment for corresponding disease indications.

Exposure	Outcome	Multivariables	OR (95% CI)	P-value
Anilide	Osteoporosis	Back pain	1.2506 (1.0589-1.4769)	0.0084
Drugs used in diabetes	Osteoporosis	Type 2 diabetes	1.1343 (1.0334-1.2450)	0.008
Thyroid preparation	Osteoporosis	Hypothyroidism	1.1434 (1.0766-1.2145)	1.30×10 ⁻⁵
Thyroid preparation	Osteoporosis with pathological fracture	Hypothyroidism	1.1735 (1.0522-1.3087)	0.004
Thyroid preparation	Drug-induced osteoporosis	Hypothyroidism	1.2739 (0.9493-1.7094)	0.1067
Immunosuppressant	Osteoporosis with pathological fracture	Rheumatoid arthritis	0.9597 (0.8486-1.0853)	0.5121
Immunosuppressant	Drug-induced osteoporosis	Rheumatoid arthritis	1.0847 (0.8458-1.3911)	0.522

Abbreviations: OR, odds ratio; CI, confidence interval.

Estimates were derived using MVMR analyses adjusting for the underlying disease traits corresponding to each medication exposure (type 2 diabetes, hypothyroidism, rheumatoid arthritis, or back pain). Odds ratios represent the effect of genetically predicted medication use on osteoporosis-related outcomes independent of these confounders.



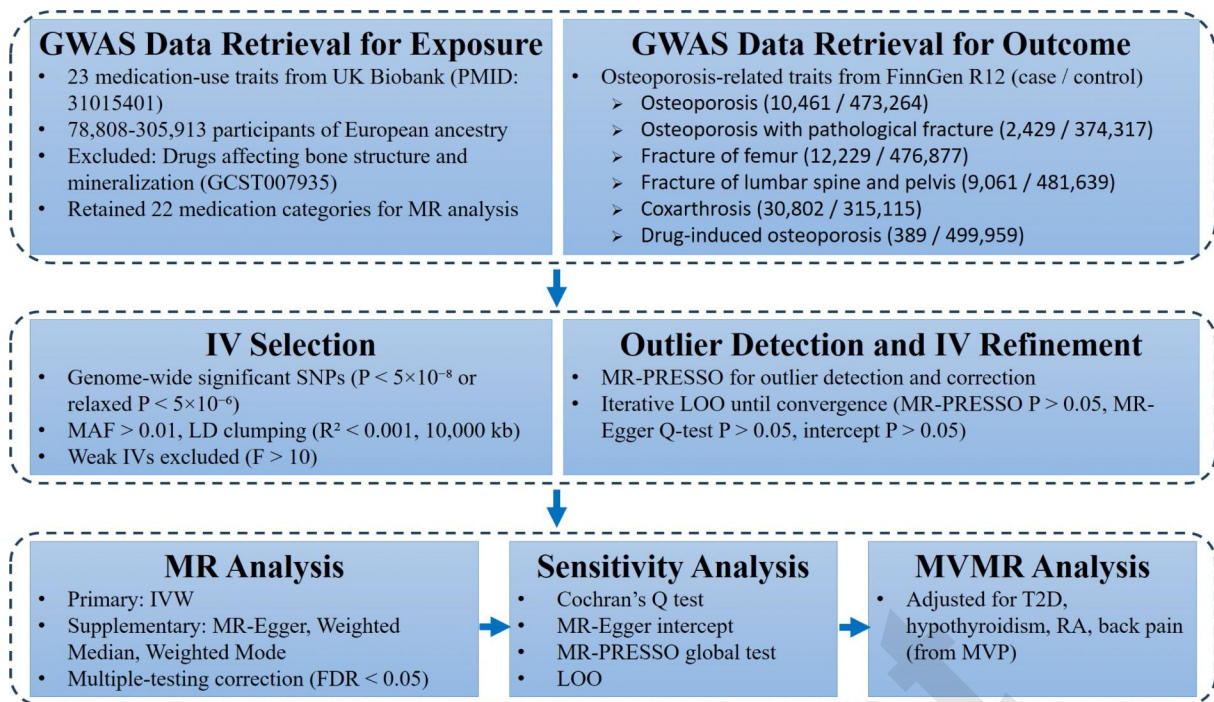


Figure 1. Overview of the study design. Summary-level data for 22 medication-use traits from the UK Biobank and osteoporosis-related outcomes from the FinnGen R12 cohort were analyzed using Mendelian randomization (MR). Independent single-nucleotide polymorphisms (SNPs) were selected as instrumental variables, followed by outlier correction, causal effect estimation, and sensitivity analyses. Multivariable MR (MCMR) was further performed to account for disease-related confounding. GWAS, genome-wide association study; MR, Mendelian randomization; MVMR, multivariable Mendelian randomization; SNP, single nucleotide polymorphism; IV, instrumental variable; IVW, inverse variance-weighted; LD, linkage disequilibrium; LOO, leave-one-out; MR-PRESSO, Mendelian Randomization Pleiotropy RESidual Sum and Outlier; T2D, type 2 diabetes; RA, rheumatoid arthritis.

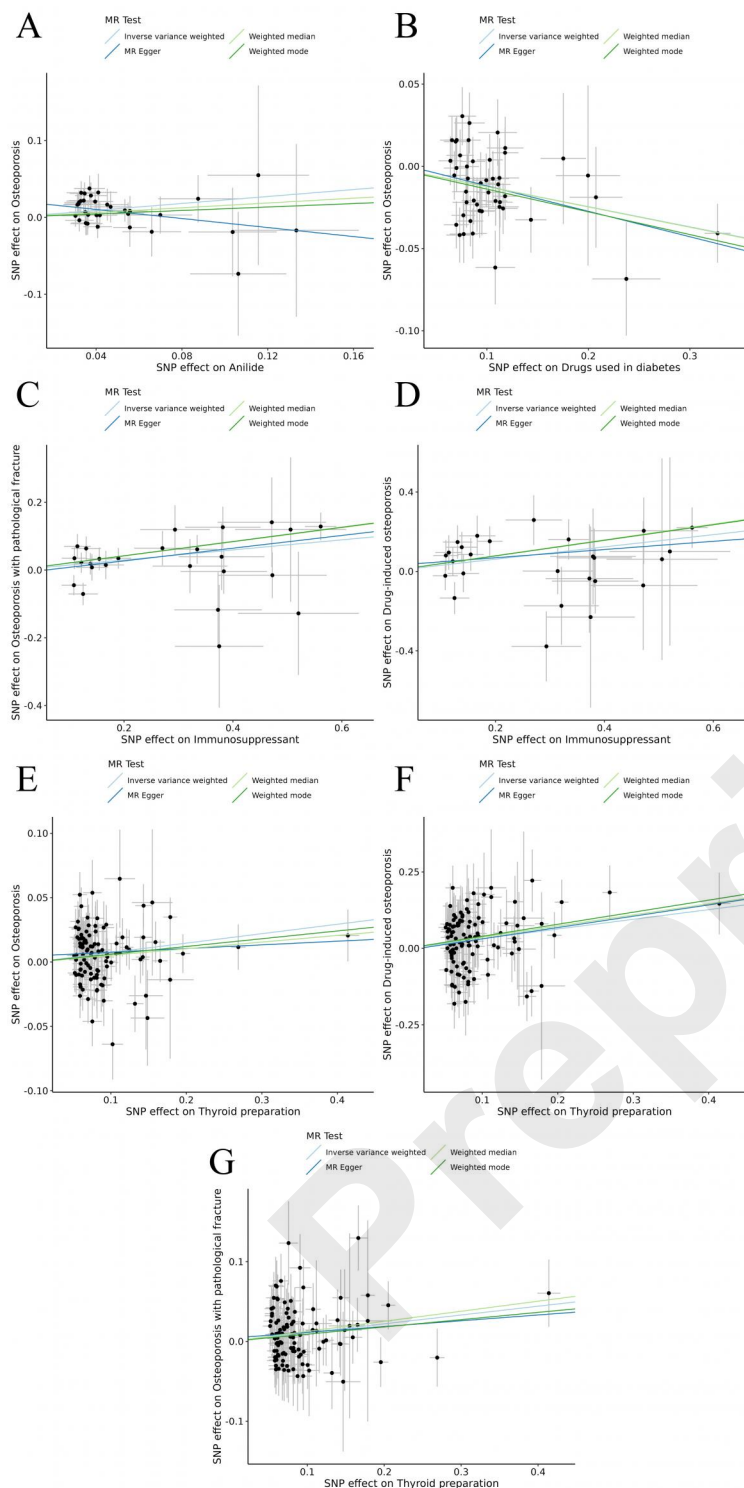


Figure 2. Scatter plots of SNP-specific causal estimates for significant associations between medication use and osteoporosis-related outcomes (FDR < 0.05). Each point represents an individual SNP, with slopes indicating causal estimates derived from different MR methods (IVW, MR-Egger, weighted median, and weighted mode). (A) Anilide use–osteoporosis, (B) Antidiabetic drug use–osteoporosis, (C) Immunosuppressant use–osteoporosis with pathological fracture, (D) Immunosuppressant use–drug-induced osteoporosis, (E) Thyroid preparation use–osteoporosis, (F) Thyroid preparation use–drug-induced osteoporosis, and (G) Thyroid preparation use–osteoporosis with pathological fracture.